

Hypothesis

Is council tax valuation band a predictor of mortality?

Norman R Beale*¹, Gordon J Taylor² and Dawn MK Straker-Cook³

Address: ¹Northlands R & D General Practice, Calne, United Kingdom, ²Bath & Swindon RDSU, Royal United Hospital, Bath, United Kingdom and ³Northlands R & D General Practice, Calne, United Kingdom

E-mail: Norman R Beale* - nbrndcal@aol.com; Gordon J Taylor - g.j.taylor@bath.ac.uk; Dawn MK Straker-Cook - nbrndcal@aol.com

*Corresponding author

Published: 3 September 2002

Received: 21 March 2002

BMC Public Health 2002, **2**:17

Accepted: 3 September 2002

This article is available from: <http://www.biomedcentral.com/1471-2458/2/17>

© 2002 Beale et al; licensee BioMed Central Ltd. This article is published in Open Access: verbatim copying and redistribution of this article are permitted in all media for any non-commercial purpose, provided this notice is preserved along with the article's original URL.

Abstract

Background: All current UK indices of socio-economic status have inherent problems, especially those used to govern resource allocation to the health sphere. The search for improved markers continues: this study proposes and tests the possibility that Council Tax Valuation Band (CTVB) might match requirements.

Presentation of the hypothesis: To determine if there is an association between CTVB of final residence and mortality risk using the death registers of a UK general practice.

Testing the hypothesis: Standardised death rates and odds ratios (ORs) for groups defined by CTVB of dwelling (A – H) were calculated using one in four denominator samples from the practice lists. Analyses were repeated three times – between number of deaths and CTVB of residence of deceased 1992 – 1994 inclusive, 1995 – 1997 inc., 1998 – 2000 inc.

In 856 deaths there were consistent and significant differences in death rates between CTVBs: above average for bands A and B residents; below average for other band residents. There were significantly higher ORs for A, B residents who were female and who died prematurely (before average group life expectancy).

Implications of the hypothesis: CTVB of final residence appears to be a proxy marker of mortality risk and could be a valuable indicator of health needs resource at household level. It is worthy of further exploration.

Background

Mortality statistics remain the least ambiguous measure of significant illness, certainly beyond youth and early middle age. Therefore, when testing the validity of any supposed predictor of health, its association with mortality will be an 'acid test' and many of the recent reports on the relationship between socio-economic status and health have used mortality data [1–17]. These late 20th century studies confirm what had been detected in the 19th century [18] – that poor socio-economic circumstances are as-

sociated with reduced life span. But our understanding of the mechanisms involved remains defective, perhaps because there is no way of relating the many different, and often proxy, measures of socio-economic status: social standing lacks the definitude of death. Researchers have used occupational role [19], levels of educational attainment [12], ad hoc combinations of occupation, home and car ownership [20], refined indices such as that of Carstairs [21] or of Townsend [1], and so-called 'geodemographic' markers [15]. There are problems with all of

these. For instance car ownership is an unavoidable expense for poor rural families and how, precisely, can one allocate occupational standing to the unemployed, the housewife, or the retired? And as often as locality comparisons are reported, critics remind us of the ecological 'fallacy' [22] viz. that 'average' endowment of an area, however small, will mask individual or household deprivation. The search for a valid, universal, current, and easily accessed UK measure of individual socio-economic standing continues, and particularly for primary care [23–27]. Resource allocation here is obviously a vital activity if limited health service provision is to be effectively targeted, an ambition reactivated by the Government's new policy of 'breaking the link between poverty and ill health' [28].

In 1992 the British Government introduced legislation [29] that mandated all UK Local Authorities to levy a new tax – the Council Tax – to replace the Community Charge. Homes were to be assessed externally, allotted an 'open market' value (as at 1 April 1991) based on size, layout, character and locality, and placed into one of 8 'Valuation Bands' A – H that would dictate the amount of the tax. In an earlier report [30] we demonstrated an association between Council Tax Valuation Band (CTVB) and:

- (i) established deprivation indicators viz. home ownership, access to a car;
- (ii) clinical workload in a typical UK general practice, i.e. recorded morbidity.

We were able to show, therefore, that CTVB matches established deprivation markers such as 'Townsend' [1] and that it also correlates (inversely) with patient demand. These associations are reinforced by our further finding [31], of a highly significant correlation between CTVBs and the Underprivileged Area 8 ('UPA8' or 'Jarman') scores [26] which are used to underpin the subsidy payments made to UK general practices in 'deprived' areas of the country. We have now extended our investigations of CTVB using mortality statistics.

Presentation of the hypothesis

The aim of this study is to test the hypothesis that the CTVB of the home in which a subject resides is a predictor of their mortality.

Testing the hypothesis

The Practice

Five whole time partners practise in equal partnership in this semi-rural Research and Development/Training practice based in modern premises. The complements of practice nurses and of administrative staff are above local averages. Patient list size has been between 10,500 & 11,000 for over a decade: age spectrum and annual turno-

ver are close to local and national averages. There are no significant ethnic minorities but Calne is, historically, an industrial town with pockets of deprivation (two enumeration districts have very high UPA8 scores) located in a sparsely populated and largely agrarian county.

Study design

The blueprint was to derive the death rates (standardised directly using our study population norms) of patients who were fully registered with the study practice – grouped by CTVB of residence – over the nine potential years of study, 1992 – 2000 inclusive.

Numerator data (all deaths)

The mortality data were retrieved from the practice archives (collected routinely by staff for our annual practice reports) and corroborated by the regular 'deduction' lists received from Wiltshire Health Authority for the return of case notes of patients no longer registered with the practice (because of death in this instance). The archived material provided full name, gender, date of birth, date of death, registered doctor and home address. Individuals living in residential or nursing homes (who are exempt from Council Tax [29]) were admitted to the study by a process of reverse attribution to their former private residential addresses. The data were then anonymised and aggregated into three triennial sets to match the 3 denominator samples i.e. deaths in 1992, 1993, & 1994, deaths in 1995, 1996, & 1997, and deaths in 1998, 1999, & 2000 respectively (see below).

Denominator data (25% samples of practice lists)

Three representative sets of denominator data were obtained from Wiltshire Health Authority records – viz. for July 1, 1993; July 1, 1996; and July 1, 1999 (to minimise unscheduled work for Health Authority staff). The data were aggregated from two sources in appropriate proportions:

- (i) computerised files provided by the Health Authority for the 'live' lists for the practice i.e. patients who had been registered at the pertinent times and who were still registered;
- (ii) archived (hard copy) 'deduction lists' from the Health Authority i.e. of patients who had been removed from the practice lists since each of the three sampling dates.

The extracted data consisted of full name, gender, date of birth, registered doctor and home address. Reverse attribution was again applied and data anonymised. To make the data manageable, randomiser methods were used to select 25% samples of the denominator populations.

Table 1: Number of deaths, n. 975, divided by age band, genders separated

Age band, years		0-9	10-19	20-29	30-39	40-49	50-59	60-69	70-79	80-89	90+	Totals
No. of deaths	males	4	2	5	4	13	29	68	144	158	32	459
	females	0	2	0	2	8	21	65	128	186	104	516
	totals	4	4	5	6	21	50	133	272	343	136	975

Table 2: Numbers of study deaths in the three study sets (over 50 years) distributed by CTVB of residence of the deceased (median age at death in parentheses)

CTVB	A	B	C	D	E	F+
1993	67 (79)	89 (79)	52 (76)	58 (79)	32 (77)	10 (81)
1996	68 (81)	99 (79)	42 (77)	34 (82)	16 (82)	15 (85)
1999	49 (77)	89 (79)	56 (81)	45 (81)	19 (78)	16 (78)

Council Tax Valuation Bands

The CTVB [A – H] for the home addresses of (a) the deceased and of (b) the denominator samples, were obtained from the Council Tax Valuation Lists published by, and available for reference at the offices of, the relevant Local Authorities [North Wiltshire and Kennet]. Banded assessments were originally made in 1991 & 1992: new or amended properties have all been subsequently assessed contemporaneously but then devalued to 1991 levels [29] for banding.

Statistical testing

The data are represented by actual numbers, standardised death rates, and odds ratios. All of the data in the analyses are from the 50 years and over practice population. The denominators are based on 25% samples of the 50 years and over practice populations: twenty five percent was chosen for logistic reasons. However, this is sufficient to provide us with stable estimates of the frequency distributions across the CTVB categories. Differences in the ages at death between the groups were tested by the Kruskal-Wallis H test. National data on population structure by CTVB group are not available: therefore the expected numbers of deaths used to calculate the standardised death rates (mean = 100) are calculated as the total number of deaths divided by the total size of the denominator population and then multiplied by the number in each CTVB group. Where appropriate, 95% confidence intervals have been calculated using the methods described in Altman (1999) [33]. Owing to the small number of deaths in these categories, aggregation of CTVBs F, G, and H was necessary: this was only undertaken in order to maintain sufficient numbers in each category (expected number of deaths not

less than 5). When considering the odds ratios, CTVB is further reduced down to two categories based on the post-hoc observation of a change in the standardised death rates from greater to less than 100.

Ethical approval was granted by the Bath Local Ethics Committee.

Numerator findings

Between 1 January 1992 and 31 December 2000 there were 975 deaths of patients registered at the practice; 459 were male (47.1%) and 516 were female (52.9%). Annual totals varied between 91 and 130 deaths (mean 108). There was no discernible time trend. The spectrum of deaths by age at death is given in Table 1. For women in the study the median age (inter-quartile range) at death is 82 (73 – 88) years contrasting to 77 (69 – 84) years for males.

There were only 40 individuals who died aged less than 50 years (4.1%) – 28 males and 12 females. All subsequent analyses (tables 2,3,4,5,6,7) were therefore restricted to numerator data (n. 935) and denominator data for those aged 50 years and above. Reverse attribution to former private address was applied to 59 Nursing Home residents and to 76 Residential Home occupants. There were insufficient data to so locate 58 of the 935 deaths (6.2%) because although dying in local Care Homes (34 at Nursing Homes, 24 at Residential Homes) they had formerly lived outside the Local Authority areas.

Otherwise there were only 21 deceased (2% of all study deaths) for whom our practice archives and those of the

Table 3: Denominator frequencies i.e. 25% of practice populations (over 50 years) by CTVB in the 3 study sets - as at 1 July 1993, 1 July 1996, 1 July 1999

	CTVB	A	B	C	D	E	F+
<i>Denominators</i>							
1993		121	209	185	175	103	87
1996		104	222	208	184	97	73
1999		87	187	196	216	142	116

Table 4: Standardised death rates for each of the three study (over 50 years) sets subdivided by CTVB of residence

1993 study set							
	CTVB	A	B	C	D	E	F+
<i>Deaths</i>		67	89	52	58	32	10
<i>Denoms</i>		121	209	185	175	103	87
SDR		158	122	80	95	88	33
95% CIs		120 – 196	96 – 147	58 – 102	70 – 119	58 – 120	12 – 53
1996 study set							
	CTVB	A	B	C	D	E	F+
<i>Deaths</i>		68	99	42	34	16	15
<i>Denoms</i>		104	222	208	184	97	73
SDR		212	145	65	60	53	67
95% CIs		162 – 262	116 – 173	46 – 85	40 – 80	27 – 79	33 – 100
1999 study set							
	CTVB	A	B	C	D	E	F+
<i>Deaths</i>		49	89	56	45	19	16
<i>Denoms</i>		87	187	196	216	142	116
SDR		194	164	98	72	46	48
95% CIs		140 – 248	130 – 198	73 – 124	51 – 93	25 – 67	24 – 71

Health Authority were jointly insufficient for unambiguous identification. The refined dataset for full analysis was therefore 856 deaths. The numbers of deaths (and median ages at death) according to determined CTVB of residence are shown in Table 2 where bands F,G,H are again united as 'F+'. There are no significant differences in age at death between the CTVB subgroups in any of the 3 study sets.

Denominator findings

25% randomised samples from the practice lists (50 years and over) as at 1 July 1993, 1996, & 1999, subdivided by CTVB of then residence (F, G, & H amalgamated because of small numbers), are shown in Table 3. We were able to locate all but 28 individuals (1.0%) to a CTVB.

Table 5: Standardised death rates for CTVB A,B residents compared with those for C+ residents with the Odds Ratios for former in relation to latter

1993 study set			
CTVB	A,B	C +	Odds Ratio A,B/C+ (95% CI)
Deaths	156	152	
Denominators	330	550	
Mortality Risk	0.473	0.276	1.71 (1.41 – 2.07)
1996 study set			
CTVB	A,B	C +	Odds Ratio A,B/C+ (95% CI)
Deaths	167	107	
Denominators	326	562	
Mortality Risk	0.512	0.190	2.69 (2.17 – 3.33)
1999 study set			
CTVB	A,B	C +	Odds Ratio A,B/C+ (95% CI)
Deaths	138	136	
Denominators	274	670	
Mortality Risk	0.504	0.203	2.48 (2.02 – 3.05)

Standardised death rates (SDRs) death and denominator findings and SDRs (95% C Is) subdivided by CTVB for the three study sets are shown in Table 4. The risk of dying during the study period is consistently higher for residents of CTVB A, B homes than for those in the higher bands viz. C and above, significantly so for the '1996' and '1999' sets.

In fact the conventional threshold of SDR 100 clearly differentiates two distinct populations – those deceased who had resided in CTVB A, B homes (SDRs always above 100) from those deceased who had resided in CTVB 'C+' homes (all SDRs below 100). This observation permitted further analyses viz. calculating Odds Ratios for A, B versus C+, again repeated in each of the three sets (see Table 5). These analyses were then extended to test for the influence of gender of the deceased in respective groups and of the age at death (below, versus above, median age at death of respective group). The results consistently show that the A, B residents are at significantly greater risk of dying than their counterparts residing in bands C and above and that socio-economic influence, as marked by CTVB, is higher in women and in those who die before median life expectancy (Tables 6,7).

Implications of the hypothesis

This is a modest study in general practice ecology. Its uses simple techniques and its conclusions should be viewed with caution. Its findings are, however, clear and consistent. In fact we have performed the study, effectively, three times and the results superimpose: there is a striking downward trend in death rates with increasing CTVB of final private residence which is not confounded by age at death. In other words those middle-aged and elderly of our practice population who live in the most modest homes are the most likely to die in any time interval: our study hypothesis is supported. These observations reinforce earlier findings from our practice [30] where we found a parallel (inverse) association between CTVB and illness: the higher the band, the lower the recorded morbidity. And mortality risk for those with death rates above average (residents of CTVB A and B) compared with their counterparts reveals the greatest differences to be in women and in those whom we might describe as dying prematurely – those who die before attaining group life expectancy. Perhaps these are all clues for where we should focus limited clinical resources.

We were unable, in this study, to account for the influence of some factors that might be construed as confounders of

Table 6: Standardised death rates for CTVB A,B residents (males and females differentiated) compared with those for C+ residents with the Odds Ratios for former in relation to latter

1993 study set		A,B	C+	Odds Ratios (A,B/C+) & 95% CIs
Males	Deaths	74	81	1.38 (1.06 – 1.81)
	Denominator	171	259	
	Mortality Risk	0.433	0.313	
Females	Deaths	82	71	2.11 (1.61 – 2.78)
	Denominator	159	291	
	Mortality Risk	0.516	0.244	
1996 study set		A,B	C+	Odds Ratios (A,B/C+) & 95% CIs
Males	Deaths	68	57	2.26 (1.67 – 3.08)
	Denominator	141	268	
	Mortality Risk	0.482	0.213	
Females	Deaths	99	50	3.15 (2.32 – 4.25)
	Denominator	185	294	
	Mortality Risk	0.535	0.170	
1999 study set		A,B	C+	Odds Ratios (A,B/C+) & 95% CIs
Males	Deaths	57	74	1.93 (1.43 – 2.60)
	Denominator	125	313	
	Mortality Risk	0.456	0.236	
Females	Deaths	81	62	3.10 (2.34 – 4.18)
	Denominator	149	357	
	Mortality Risk	0.54	0.174	

our findings: widowhood and living alone are two obvious examples. On the other hand it might be that CTVB represents a categorical and valuable proxy marker of these known disadvantages giving social scientists and public health physicians an indication of where to focus preventative health measures. And the spectrum of the CTVBs of all the patients registered with a general practitioner could provide a unique predictor of the likely morbidity and mortality levels of that practice: again resources could be matched to true need, the only real answer to health inequality.

CTVB may eventually prove to be a reliable and generic marker of socio-economic status in the UK. It certainly has many inherent strengths, being assessed independently of health and social agencies, universally attributed, officially recorded, discriminating to individual household level, and being in the public domain (now on the internet [34]). It is also free of the 'ecological fallacy' [22] and of other drawbacks of Census data such as skewed under-

enumeration [35]. However, our findings need corroboration by other, larger studies before we speculate further.

Competing interests

None declared.

Authors' contributions

NRB & DMKS-C designed the study, performed the data collection and wrote the paper. GJT performed the statistical analysis. NRB conceived the study.

All authors read and approved the final manuscript.

Acknowledgements

We are grateful for the help of the following: Professor George Davey-Smith; Professor Graham Watt; Dr. Noel Olsen; Mr. T. Brown, Council Tax Valuation Lists manager, Wessex; Liam Scanlon, Sarah Hartley, Jim Sinclair, Kath Leach and Maureen Dobson (of Wiltshire Health Authority); all partners and staff at Northlands Surgery; Chris Moreton and Jay Montague of Wootten's, Funeral Directors, Calne. Northlands Research and Development activities are supported by a Grant from the NHSE.

Table 7: Standardised death rates for CTVB A,B residents (older and younger deceased differentiated) compared with those for C+ residents with the Odds Ratios for former in relation to latter

1993 study set		A,B	C+	Odds Ratio (A,B/C+) & 95% CIs
Below median age at death	Deaths	71	84	
	Denominator	271	491	
	Mortality Risk	0.262	0.171	1.53 (1.15 – 2.04)
At or above median age at death	Deaths	85	68	
	Denominator	59	59	
	Mortality Risk	1.441	1.153	1.25 (1.01 – 1.54)
1996 study set		A,B	C+	Odds Ratio (A,B/C+) & 95% CIs
Below median age at death	Deaths	83	52	
	Denominator	285	504	
	Mortality Risk	0.291	0.103	2.83 (2.05 – 3.88)
At or above median age at death	Deaths	84	55	
	Denominator	41	58	
	Mortality Risk	2.048	0.948	2.16 (1.72 – 2.71)
1999 study set		A,B	C+	Odds Ratio (A,B/C+) & 95% CIs
Below median age at death	Deaths	70	62	
	Denominator	228	606	
	Mortality Risk	0.307	0.102	3.01 (2.19 – 4.11)
At or above median age at death	Deaths	68	74	
	Denominator	47	66	
	Mortality Risk	1.447	1.12	1.29 (1.04 – 1.60)

References

- Townsend P, Phillimore P, Beattie A: **Health and deprivation: inequality and the north.** *Croom Helm, Beckenham, Kent* 1988
- Townsend P, Davidson N, Whitehead M, eds: **Inequalities in Health. The Black Report and The Health Divide.** *Penguin Books, Harmondsworth, Middlesex* 1992
- Marmot M, Davey Smith G, Stansfield S, Patel C, et al: **Health inequalities among British civil servants: the Whitehall II study.** *The Lancet* 1991, **337**:1387-1393
- Sloggett A, Joshi H: **Higher mortality in deprived areas: community or personal disadvantage?** *BMJ* 1994, **309**:1470-1474
- Borrell C, Arias A: **Socioeconomic factors and mortality in urban settings: the case of Barcelona, Spain.** *J Epidemiol Comm Health* 1995, **49**:460-465
- Duncan B, Rumel D, Zelmanowicz A, Mengue S, et al: **Social inequality in mortality in Sao Paulo State, Brazil.** *Int J of Epidemiol* 1995, **24**:359-365
- Bennett S: **Socioeconomic inequalities in coronary heart disease and stroke mortality among Australian men 1979–1993.** *Int J Epidemiol* 1996, **25**:266-275
- Jessop E: **Deprivation and mortality in non-metropolitan areas of England and Wales.** *J Epidemiol & Comm Health* 1996, **50**:524-526
- Kaplan G, Pamuk E, Lynch J, Cohen R, Balfour J: **Inequality in income and mortality in the United States: analysis of mortality and potential pathways.** *BMJ* 1996, **312**:999-1003
- Kennedy B, Kawachi I, Prothrow-Stith D: **Income distribution and mortality: cross sectional ecological study of the Robin Hood index in the United States.** *BMJ* 1996, **312**:1004-1007
- Marmot M, Shipley M: **Do socio-economic differences in mortality persist after retirement? 25 year follow up of civil servants from the first Whitehall study.** *BMJ* 1996, **313**:1177-1180
- Morris J, Blane D, White I: **Levels of mortality, education, and social conditions in the 107 local education authority areas of England.** *J Epidemiol Comm. Health* 1996, **50**:15-17
- Davey Smith G, Hart C, Blane D, Gillis G, Hawthorne V: **Lifetime socio-economic position and mortality: prospective observational study.** *BMJ* 1997, **314**:547-552
- Wannamathee S, Shaper G: **Socioeconomic status within social class and mortality: a prospective study in middle-aged British men.** *Int. J. Epidemiol* 1997, **26**:532-541

15. Dorling D: **Death in Britain. How local mortality rates have changed: 1950s – 1990s.** *Joseph Rowntree Foundation, York* 1997
16. Maheswaran R, Elliott P, Strachan D: **Socioeconomic deprivation, ethnicity, and stroke mortality in Greater London and south east England.** *J Epidemiol & Comm Health* 1997, **51**:127-131
17. Shaw M, Dorling D, Gordon D, Davey Smith G: **The widening gap.** *The Policy Press, Bristol* 1999
18. Rowntree BB: **Poverty: a study of town life.** *MacMillan, London* 1901
19. Marmot M, Shipley M, Rose G: **Inequalities in death – specific explanations of a general pattern?** *The Lancet* 1984, **I**:1003-1006
20. Filakti H, Fox J: **Differences in mortality by housing tenure and by car access from the OPCS Longitudinal Study.** *Population Trends* 1995, **81**:27-30
21. Carstairs V, Morris R: **Deprivation and health in Scotland.** *Aberdeen University Press, Aberdeen* 1991
22. Robinson W: **Ecological correlations and the behaviour of individuals.** *American Sociological Reviews* 1950, **15**:351-357
23. Davey Smith G, Dorling D: **"I'm all right, John": voting patterns and mortality in England and Wales, 1981 – 1992.** *BMJ* 1996, **313**:1573-1577
24. Lloyd C, Harris C, Clucas D: **Low income scheme index: a new deprivation scale based on prescribing in general practice.** *BMJ* 1995, **310**:165-170
25. Campbell D, Radford J, Burton P: **Unemployment rates; an alternative to the Jarman index?** *BMJ* 1991, **303**:750-755
26. Jarman B: **Identification of underprivileged areas.** *BMJ* 1983, **286**:1705-1709
27. Beale N: **Unequal to the task: deprivation, health and UK general practice at the millennium.** *British Journal of General Practice* 2001, **51**:478-485
28. Anonymous: **The National Health Inequalities Targets.** [<http://www.doh.gov.uk/healthinequalities>]
29. Anonymous: **Local Government Finance Act 1992.** *Council Tax. London: HMSO* 1992
30. Beale N, Baker N, Straker-Cook D: **Council tax valuation band as marker of deprivation and of general practice workload.** *Public Health* 2000, **114**:260-264
31. Beale N, Taylor G, Straker-Cook D: **Does Council Tax Valuation Band (CTVB) correlate with Under-Privileged Area 8 (UPA8) score and could it be a better 'Jarman Index'?** *BMC Public Health*: [<http://www.biomedcentral.com/1471-2458/1/13>]
32. Armitage P: **Statistical methods in medical research.** *Blackwell, Oxford* 1974
33. Altman P: **Practical statistics for medical research.** *Chapman & Hall/CRC, London* 1999
34. [www.voa.gov.uk/council_tax/index.htm]
35. Majeed F, Cook D, Poloniecki J, Martin DD: **Using data from the 1991 Census.** *BMJ* 1995, **310**:1511-1514

Pre-publication history

The pre-publication history for this paper can be accessed here:

<http://www.biomedcentral.com/1471-2458/2/17/prepub>

Publish with **BioMed Central** and every scientist can read your work free of charge

"BioMedCentral will be the most significant development for disseminating the results of biomedical research in our lifetime."

Paul Nurse, Director-General, Imperial Cancer Research Fund

Publish with **BMC** and your research papers will be:

- available free of charge to the entire biomedical community
- peer reviewed and published immediately upon acceptance
- cited in PubMed and archived on PubMed Central
- yours - you keep the copyright

Submit your manuscript here:

<http://www.biomedcentral.com/manuscript/>



BioMedcentral.com

editorial@biomedcentral.com